

Today

Finish Linear Regression:

Today

Finish Linear Regression:

Best linear function prediction of Y given X .

Today

Finish Linear Regression:

Best linear function prediction of Y given X .

MMSE: Best Function that predicts Y from S .

Today

Finish Linear Regression:

Best linear function prediction of Y given X .

MMSE: Best Function that predicts Y from S .

Conditional Expectation.

Today

Finish Linear Regression:

Best linear function prediction of Y given X .

MMSE: Best Function that predicts Y from S .

Conditional Expectation.

Applications to random processes.

LLSE

LLSE

Theorem

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0,$

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra.

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β ,

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - a - bX = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$.

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - a - bX = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$.

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$E[(Y - \alpha - \beta X)^2] = E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2]$$

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \end{aligned}$$

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X])$. $E[Y - \hat{Y}] = 0$ by linearity.

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - a - bX = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$. Now,

$$\begin{aligned} E[(Y - a - bX)^2] &= E[(Y - \hat{Y} + \hat{Y} - a - bX)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - a - bX)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that $E[(Y - \hat{Y})^2] \leq E[(Y - a - bX)^2]$, for all (a, b) .

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X])$. $E[Y - \hat{Y}] = 0$ by linearity.

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - a - bX = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$. Now,

$$\begin{aligned} E[(Y - a - bX)^2] &= E[(Y - \hat{Y} + \hat{Y} - a - bX)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - a - bX)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that $E[(Y - \hat{Y})^2] \leq E[(Y - a - bX)^2]$, for all (a, b) .

Thus \hat{Y} is the LLSE.

LLSE

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.
Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$. $E[Y - \hat{Y}] = 0$ by linearity.

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (See next slide.)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that $E[(Y - \hat{Y})^2] \leq E[(Y - a - bX)^2]$, for all (a, b) .

Thus \hat{Y} is the LLSE. □

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]} (X - E[X]).$$

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$.

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because $E[(Y - \hat{Y})E[X]] = 0$.

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because $E[(Y - \hat{Y})E[X]] = 0$.

Now,

$$\begin{aligned} & E[(Y - \hat{Y})(X - E[X])] \\ &= E[(Y - E[Y])(X - E[X])] - \frac{\text{cov}(X, Y)}{\text{var}[X]} E[(X - E[X])(X - E[X])] \end{aligned}$$

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because $E[(Y - \hat{Y})E[X]] = 0$.

Now,

$$\begin{aligned} E[(Y - \hat{Y})(X - E[X])] &= E[(Y - E[Y])(X - E[X])] - \frac{\text{cov}(X, Y)}{\text{var}[X]} E[(X - E[X])(X - E[X])] \\ &=^{(*)} \text{cov}(X, Y) - \frac{\text{cov}(X, Y)}{\text{var}[X]} \text{var}[X] = 0. \quad \square \end{aligned}$$

(*) Recall that $\text{cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$ and $\text{var}[X] = E[(X - E[X])^2]$.

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Or what is the mean squared estimation error?

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Or what is the mean squared estimation error?

We find

$$E[|Y - L[Y|X]|^2] = E[(Y - E[Y] - (\text{cov}(X, Y)/\text{var}(X))(X - E[X]))^2]$$

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Or what is the mean squared estimation error?

We find

$$\begin{aligned} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (\text{cov}(X, Y)/\text{var}(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(\text{cov}(X, Y)/\text{var}(X))E[(Y - E[Y])(X - E[X])] \\ &\quad + (\text{cov}(X, Y)/\text{var}(X))^2 E[(X - E[X])^2] \end{aligned}$$

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Or what is the mean squared estimation error?

We find

$$\begin{aligned} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (\text{cov}(X, Y)/\text{var}(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(\text{cov}(X, Y)/\text{var}(X))E[(Y - E[Y])(X - E[X])] \\ &\quad + (\text{cov}(X, Y)/\text{var}(X))^2 E[(X - E[X])^2] \\ &= \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}. \end{aligned}$$

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Or what is the mean squared estimation error?

We find

$$\begin{aligned} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (\text{cov}(X, Y)/\text{var}(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(\text{cov}(X, Y)/\text{var}(X))E[(Y - E[Y])(X - E[X])] \\ &\quad + (\text{cov}(X, Y)/\text{var}(X))^2 E[(X - E[X])^2] \\ &= \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}. \end{aligned}$$

Without observations, the estimate is $E[Y]$.

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Or what is the mean squared estimation error?

We find

$$\begin{aligned} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (\text{cov}(X, Y)/\text{var}(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(\text{cov}(X, Y)/\text{var}(X))E[(Y - E[Y])(X - E[X])] \\ &\quad + (\text{cov}(X, Y)/\text{var}(X))^2 E[(X - E[X])^2] \\ &= \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}. \end{aligned}$$

Without observations, the estimate is $E[Y]$. The error is $\text{var}(Y)$.

Estimation Error

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

How good is this estimator?

Or what is the mean squared estimation error?

We find

$$\begin{aligned} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (\text{cov}(X, Y)/\text{var}(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(\text{cov}(X, Y)/\text{var}(X))E[(Y - E[Y])(X - E[X])] \\ &\quad + (\text{cov}(X, Y)/\text{var}(X))^2 E[(X - E[X])^2] \\ &= \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}. \end{aligned}$$

Without observations, the estimate is $E[Y]$. The error is $\text{var}(Y)$. Observing X reduces the error.

Estimation Error: A Picture

We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$$

Estimation Error: A Picture

We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$$

and

$$E[|Y - L[Y|X]|^2] = \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}.$$

Estimation Error: A Picture

We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$$

and

$$E[|Y - L[Y|X]|^2] = \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}.$$

Here is a picture when $E[X] = 0, E[Y] = 0$:

Estimation Error: A Picture

We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$$

and

$$E[|Y - L[Y|X]|^2] = \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}.$$

Here is a picture when $E[X] = 0, E[Y] = 0$:

Dimensions correspond to sample points, uniform sample space.

Estimation Error: A Picture

We saw that

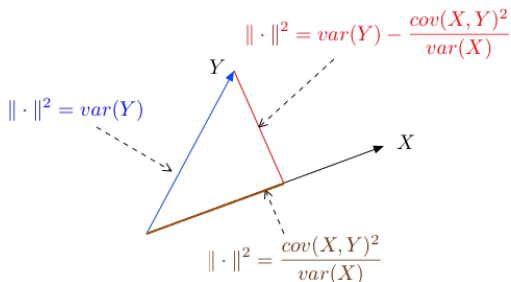
$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$$

and

$$E[|Y - L[Y|X]|^2] = \text{var}(Y) - \frac{\text{cov}(X, Y)^2}{\text{var}(X)}.$$

Here is a picture when $E[X] = 0, E[Y] = 0$:

Dimensions correspond to sample points, uniform sample space.



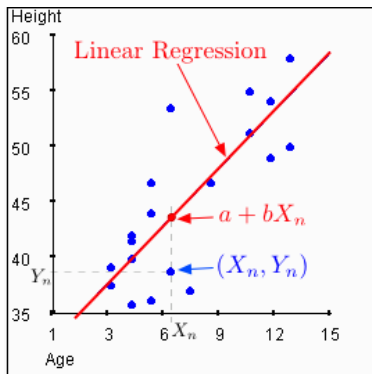
Vector Y at dimension ω is $\frac{1}{\sqrt{\Omega}} Y(\omega)$

Linear Regression Examples

Example 1:

Linear Regression Examples

Example 1:

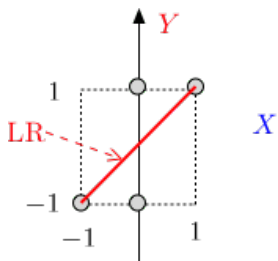


Linear Regression Examples

Example 2:

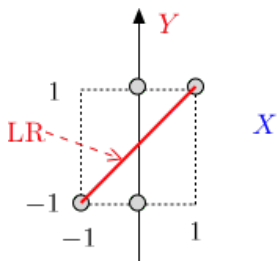
Linear Regression Examples

Example 2:



Linear Regression Examples

Example 2:

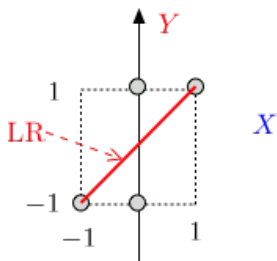


We find:

$$E[X] =$$

Linear Regression Examples

Example 2:

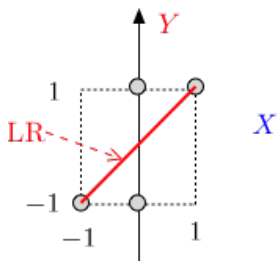


We find:

$$E[X] = 0;$$

Linear Regression Examples

Example 2:

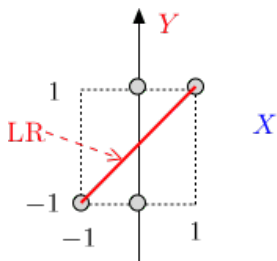


We find:

$$E[X] = 0; E[Y] =$$

Linear Regression Examples

Example 2:

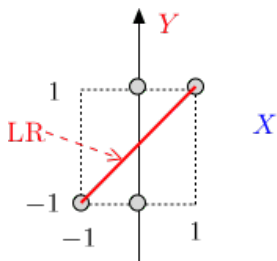


We find:

$$E[X] = 0; E[Y] = 0;$$

Linear Regression Examples

Example 2:

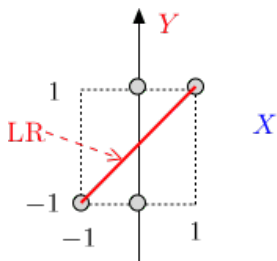


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] =$$

Linear Regression Examples

Example 2:

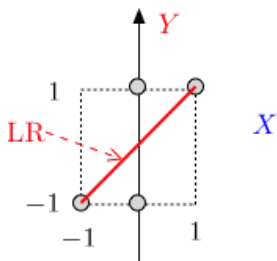


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2;$$

Linear Regression Examples

Example 2:

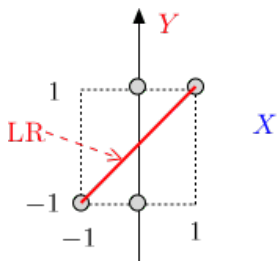


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] =$$

Linear Regression Examples

Example 2:

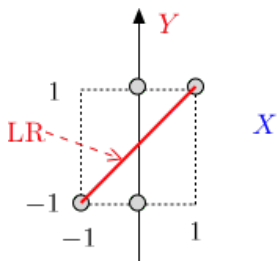


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

Linear Regression Examples

Example 2:



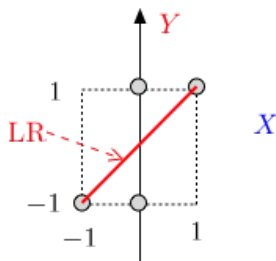
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 =$$

Linear Regression Examples

Example 2:



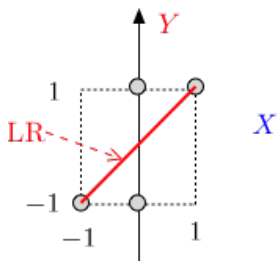
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2;$$

Linear Regression Examples

Example 2:



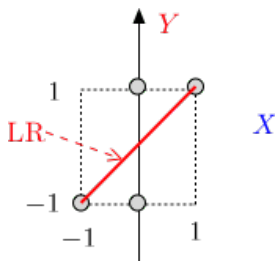
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] =$$

Linear Regression Examples

Example 2:



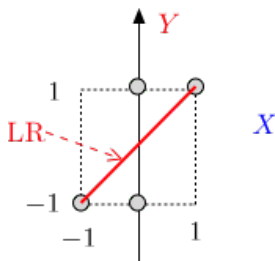
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = 1/2;$$

Linear Regression Examples

Example 2:



We find:

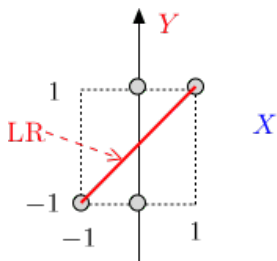
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = 1/2;$$

$$\text{LR: } \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]) =$$

Linear Regression Examples

Example 2:



We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = 1/2;$$

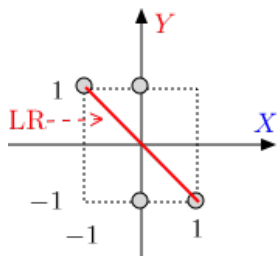
$$\text{LR: } \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]) = X.$$

Linear Regression Examples

Example 3:

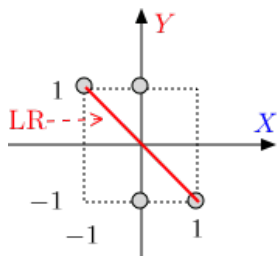
Linear Regression Examples

Example 3:



Linear Regression Examples

Example 3:

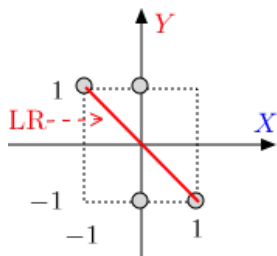


We find:

$$E[X] =$$

Linear Regression Examples

Example 3:

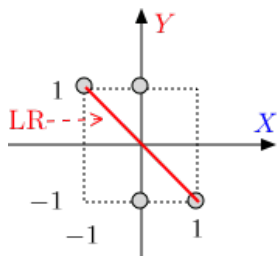


We find:

$$E[X] = 0;$$

Linear Regression Examples

Example 3:

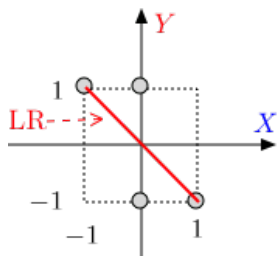


We find:

$$E[X] = 0; E[Y] =$$

Linear Regression Examples

Example 3:

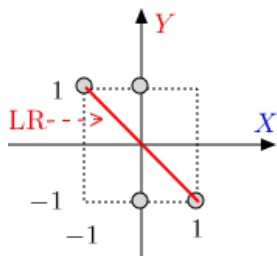


We find:

$$E[X] = 0; E[Y] = 0;$$

Linear Regression Examples

Example 3:

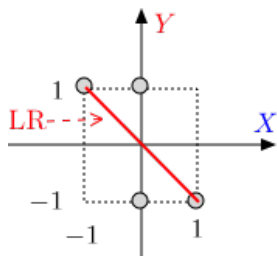


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] =$$

Linear Regression Examples

Example 3:

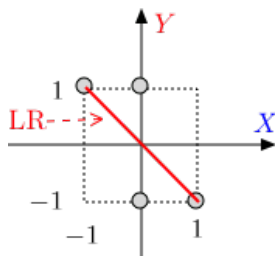


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2;$$

Linear Regression Examples

Example 3:

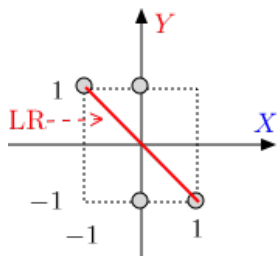


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] =$$

Linear Regression Examples

Example 3:

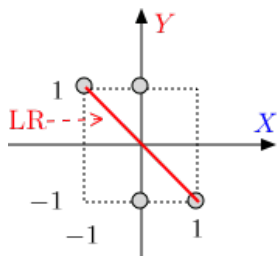


We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

Linear Regression Examples

Example 3:



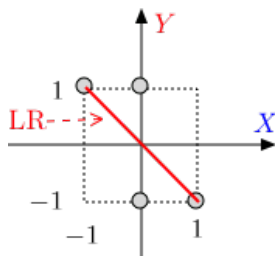
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 =$$

Linear Regression Examples

Example 3:



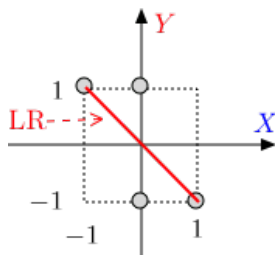
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2;$$

Linear Regression Examples

Example 3:



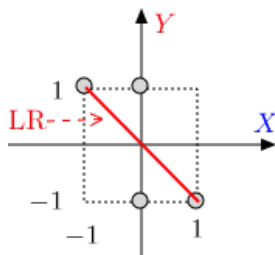
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] =$$

Linear Regression Examples

Example 3:



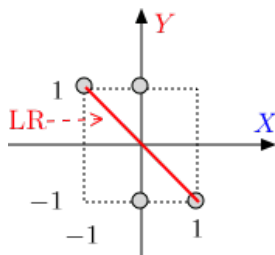
We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = -1/2;$$

Linear Regression Examples

Example 3:



We find:

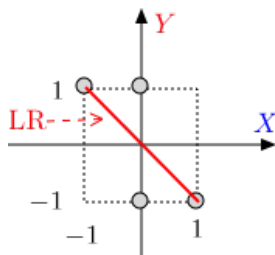
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = -1/2;$$

$$\text{LR: } \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]) =$$

Linear Regression Examples

Example 3:



We find:

$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = -1/2;$$

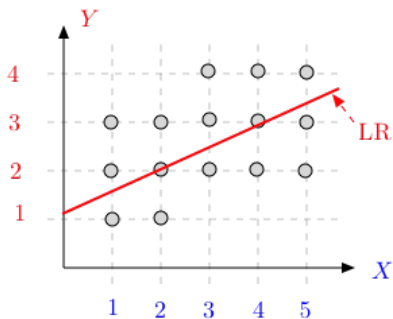
$$\text{LR: } \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]) = -X.$$

Linear Regression Examples

Example 4:

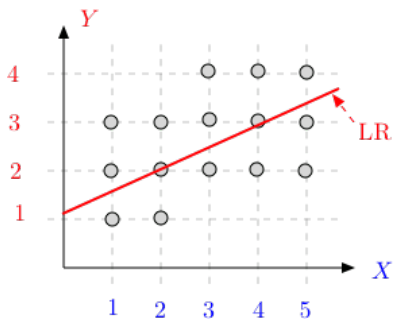
Linear Regression Examples

Example 4:



Linear Regression Examples

Example 4:

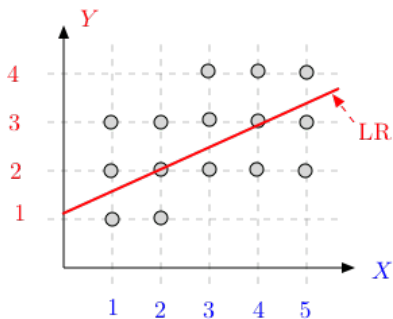


We find:

$$E[X] =$$

Linear Regression Examples

Example 4:

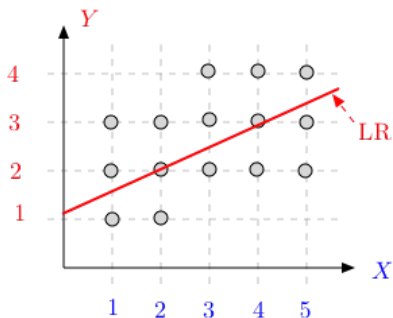


We find:

$$E[X] = 3;$$

Linear Regression Examples

Example 4:

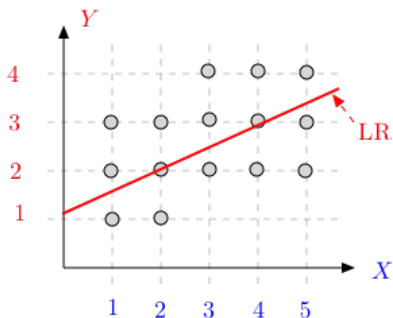


We find:

$$E[X] = 3; E[Y] =$$

Linear Regression Examples

Example 4:

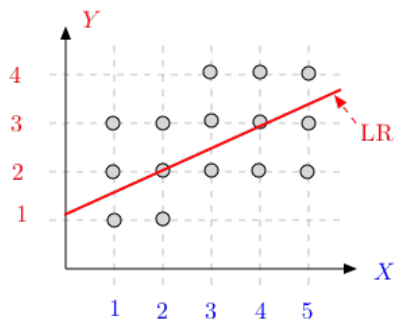


We find:

$$E[X] = 3; E[Y] = 2.5;$$

Linear Regression Examples

Example 4:

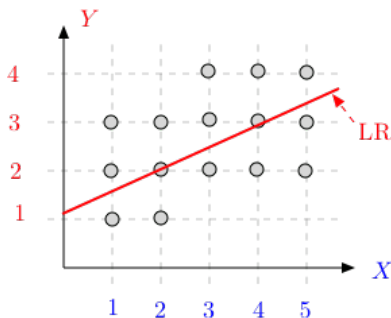


We find:

$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$

Linear Regression Examples

Example 4:



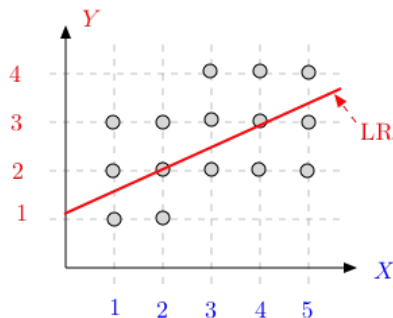
We find:

$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$

$$E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$$

Linear Regression Examples

Example 4:



We find:

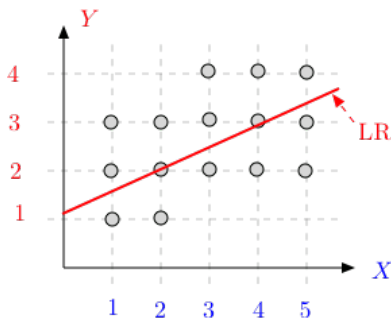
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$

$$E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$$

$$\text{var}[X] = 11 - 9 = 2;$$

Linear Regression Examples

Example 4:



We find:

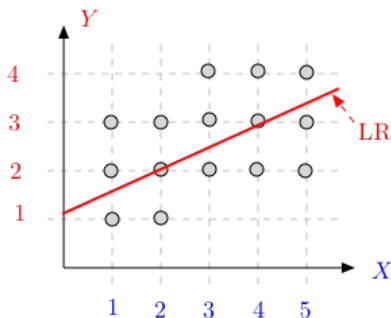
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$

$$E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$$

$$\text{var}[X] = 11 - 9 = 2; \text{cov}(X, Y) = 8.4 - 3 \times 2.5 = 0.9;$$

Linear Regression Examples

Example 4:



We find:

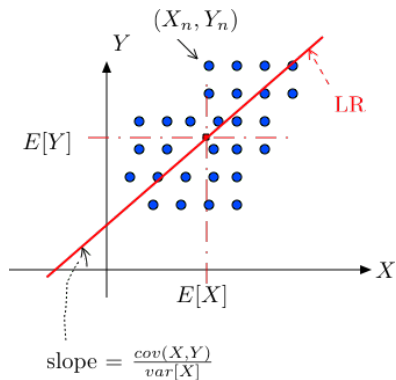
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$

$$E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$$

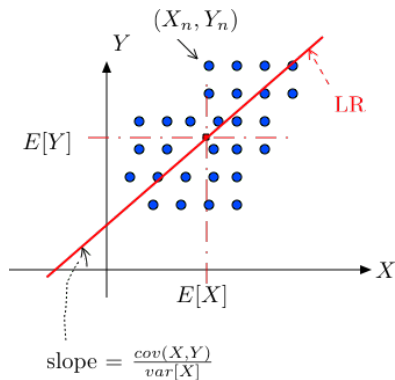
$$\text{var}[X] = 11 - 9 = 2; \text{cov}(X, Y) = 8.4 - 3 \times 2.5 = 0.9;$$

$$\text{LR: } \hat{Y} = 2.5 + \frac{0.9}{2}(X - 3) = 1.15 + 0.45X.$$

LR: Another Figure



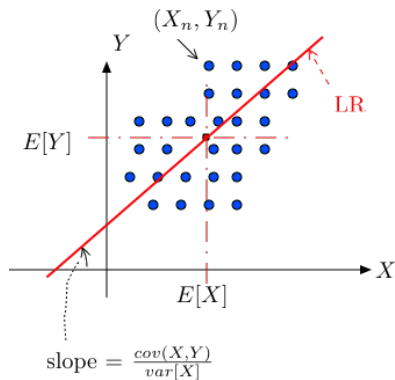
LR: Another Figure



Note that

- ▶ the LR line goes through $(E[X], E[Y])$

LR: Another Figure



Note that

- ▶ the LR line goes through $(E[X], E[Y])$
- ▶ its slope is $\frac{\text{cov}(X,Y)}{\text{var}(X)}$.

Summary

Linear Regression

Summary

Linear Regression

1. Linear Regression: $L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$

Summary

Linear Regression

1. Linear Regression: $L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$
2. Non-Bayesian: minimize $\sum_n (Y_n - a - bX_n)^2$

Summary

Linear Regression

1. Linear Regression: $L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$
2. Non-Bayesian: minimize $\sum_n (Y_n - a - bX_n)^2$
3. Bayesian: minimize $E[(Y - a - bX)^2]$

CS70: Nonlinear Regression.

1. Review: joint distribution, LLSE
2. Quadratic Regression
3. Definition of Conditional expectation
4. Properties of CE
5. Applications: Diluting, Mixing, Rumors
6. CE = MMSE

Review

Review

Definitions Let X and Y be RVs on Ω .

Review

Definitions Let X and Y be RVs on Ω .

- ▶ **Joint Distribution:** $Pr[X = x, Y = y]$

Review

Definitions Let X and Y be RVs on Ω .

- ▶ **Joint Distribution:** $Pr[X = x, Y = y]$
- ▶ **Marginal Distribution:** $Pr[X = x] = \sum_y Pr[X = x, Y = y]$

Review

Definitions Let X and Y be RVs on Ω .

- ▶ **Joint Distribution:** $Pr[X = x, Y = y]$
- ▶ **Marginal Distribution:** $Pr[X = x] = \sum_y Pr[X = x, Y = y]$
- ▶ **Conditional Distribution:** $Pr[Y = y|X = x] = \frac{Pr[X=x, Y=y]}{Pr[X=x]}$

Review

Definitions Let X and Y be RVs on Ω .

- ▶ **Joint Distribution:** $Pr[X = x, Y = y]$
- ▶ **Marginal Distribution:** $Pr[X = x] = \sum_y Pr[X = x, Y = y]$
- ▶ **Conditional Distribution:** $Pr[Y = y|X = x] = \frac{Pr[X=x, Y=y]}{Pr[X=x]}$
- ▶ **LLSE:** $L[Y|X] = a + bX$ where a, b minimize $E[(Y - a - bX)^2]$.

Review

Definitions Let X and Y be RVs on Ω .

- ▶ **Joint Distribution:** $Pr[X = x, Y = y]$
- ▶ **Marginal Distribution:** $Pr[X = x] = \sum_y Pr[X = x, Y = y]$
- ▶ **Conditional Distribution:** $Pr[Y = y|X = x] = \frac{Pr[X=x, Y=y]}{Pr[X=x]}$
- ▶ **LLSE:** $L[Y|X] = a + bX$ where a, b minimize $E[(Y - a - bX)^2]$.

We saw that

$$L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Review

Definitions Let X and Y be RVs on Ω .

- ▶ **Joint Distribution:** $Pr[X = x, Y = y]$
- ▶ **Marginal Distribution:** $Pr[X = x] = \sum_y Pr[X = x, Y = y]$
- ▶ **Conditional Distribution:** $Pr[Y = y|X = x] = \frac{Pr[X=x, Y=y]}{Pr[X=x]}$
- ▶ **LLSE:** $L[Y|X] = a + bX$ where a, b minimize $E[(Y - a - bX)^2]$.

We saw that

$$L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Recall the non-Bayesian and Bayesian viewpoints.

Nonlinear Regression: Motivation

Nonlinear Regression: Motivation

There are many situations where a good guess about Y given X is not linear.

Nonlinear Regression: Motivation

There are many situations where a good guess about Y given X is not linear.

E.g., (diameter of object, weight),

Nonlinear Regression: Motivation

There are many situations where a good guess about Y given X is not linear.

E.g., (diameter of object, weight), (school years, income),

Nonlinear Regression: Motivation

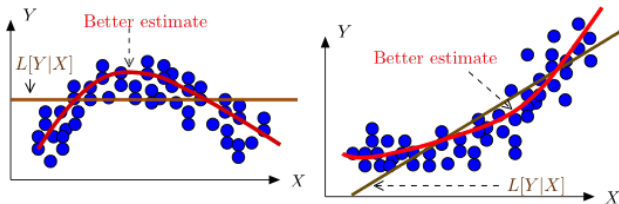
There are many situations where a good guess about Y given X is not linear.

E.g., (diameter of object, weight), (school years, income), (PSA level, cancer risk).

Nonlinear Regression: Motivation

There are many situations where a good guess about Y given X is not linear.

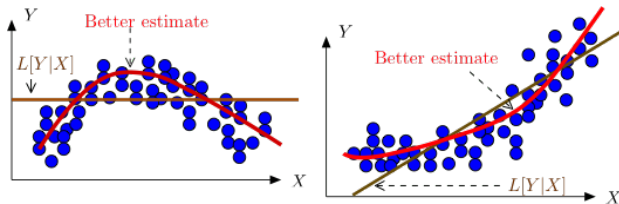
E.g., (diameter of object, weight), (school years, income), (PSA level, cancer risk).



Nonlinear Regression: Motivation

There are many situations where a good guess about Y given X is not linear.

E.g., (diameter of object, weight), (school years, income), (PSA level, cancer risk).

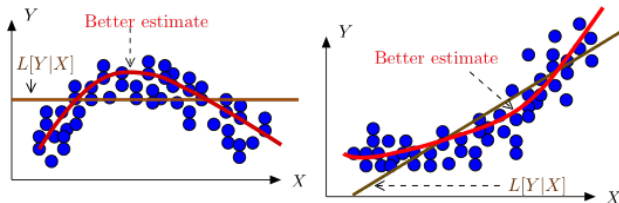


Our goal:

Nonlinear Regression: Motivation

There are many situations where a good guess about Y given X is not linear.

E.g., (diameter of object, weight), (school years, income), (PSA level, cancer risk).



Our goal: explore estimates $\hat{Y} = g(X)$ for nonlinear functions $g(\cdot)$.

Quadratic Regression

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition:

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation:

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c .

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c . We get

$$0 = E[Y - a - bX - cX^2]$$

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c . We get

$$\begin{aligned} 0 &= E[Y - a - bX - cX^2] \\ 0 &= E[(Y - a - bX - cX^2)X] \end{aligned}$$

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c . We get

$$0 = E[Y - a - bX - cX^2]$$

$$0 = E[(Y - a - bX - cX^2)X]$$

$$0 = E[(Y - a - bX - cX^2)X^2]$$

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c . We get

$$0 = E[Y - a - bX - cX^2]$$

$$0 = E[(Y - a - bX - cX^2)X]$$

$$0 = E[(Y - a - bX - cX^2)X^2]$$

We solve these three equations in the three unknowns (a, b, c) .

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c . We get

$$0 = E[Y - a - bX - cX^2]$$

$$0 = E[(Y - a - bX - cX^2)X]$$

$$0 = E[(Y - a - bX - cX^2)X^2]$$

We solve these three equations in the three unknowns (a, b, c) .

Note: These equations imply that $E[(Y - Q[Y|X])h(X)] = 0$ for any $h(X) = d + eX + fX^2$.

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c . We get

$$0 = E[Y - a - bX - cX^2]$$

$$0 = E[(Y - a - bX - cX^2)X]$$

$$0 = E[(Y - a - bX - cX^2)X^2]$$

We solve these three equations in the three unknowns (a, b, c) .

Note: These equations imply that $E[(Y - Q[Y|X])h(X)] = 0$ for any $h(X) = d + eX + fX^2$. That is, the estimation error is orthogonal to all the quadratic functions of X .

Quadratic Regression

Let X, Y be two random variables defined on the same probability space.

Definition: The quadratic regression of Y over X is the random variable

$$Q[Y|X] = a + bX + cX^2$$

where a, b, c are chosen to minimize $E[(Y - a - bX - cX^2)^2]$.

Derivation: We set to zero the derivatives w.r.t. a, b, c . We get

$$0 = E[Y - a - bX - cX^2]$$

$$0 = E[(Y - a - bX - cX^2)X]$$

$$0 = E[(Y - a - bX - cX^2)X^2]$$

We solve these three equations in the three unknowns (a, b, c) .

Note: These equations imply that $E[(Y - Q[Y|X])h(X)] = 0$ for any $h(X) = d + eX + fX^2$. That is, the estimation error is orthogonal to all the quadratic functions of X . Hence, $Q[Y|X]$ is the projection of Y onto the space of quadratic functions of X .

Conditional Expectation

Definition Let X and Y be RVs on Ω .

Conditional Expectation

Definition Let X and Y be RVs on Ω . The **conditional expectation** of Y given X is defined as

$$E[Y|X] = g(X)$$

Conditional Expectation

Definition Let X and Y be RVs on Ω . The **conditional expectation** of Y given X is defined as

$$E[Y|X] = g(X)$$

where

$$g(x) := E[Y|X = x] := \sum_y y \Pr[Y = y|X = x].$$

Conditional Expectation

Definition Let X and Y be RVs on Ω . The **conditional expectation** of Y given X is defined as

$$E[Y|X] = g(X)$$

where

$$g(x) := E[Y|X = x] := \sum_y y Pr[Y = y|X = x].$$

Fact

$$E[Y|X = x] = \sum_{\omega} Y(\omega) Pr[\omega|X = x].$$

Conditional Expectation

Definition Let X and Y be RVs on Ω . The **conditional expectation** of Y given X is defined as

$$E[Y|X] = g(X)$$

where

$$g(x) := E[Y|X = x] := \sum_y y Pr[Y = y|X = x].$$

Fact

$$E[Y|X = x] = \sum_{\omega} Y(\omega) Pr[\omega|X = x].$$

Proof: $E[Y|X = x] = E[Y|A]$ with $A = \{\omega : X(\omega) = x\}$. □

Deja vu, all over again?

Have we seen this before?

Deja vu, all over again?

Have we seen this before? Yes.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new?

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal?

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite!

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite! Simple but most convenient.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite! Simple but most convenient.

Recall that $L[Y|X] = a + bX$ is a function of X .

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite! Simple but most convenient.

Recall that $L[Y|X] = a + bX$ is a function of X .

This is similar: $E[Y|X] = g(X)$ for some function $g(\cdot)$.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite! Simple but most convenient.

Recall that $L[Y|X] = a + bX$ is a function of X .

This is similar: $E[Y|X] = g(X)$ for some function $g(\cdot)$.

In general, $g(X)$ is not linear, i.e., not $a + bX$.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite! Simple but most convenient.

Recall that $L[Y|X] = a + bX$ is a function of X .

This is similar: $E[Y|X] = g(X)$ for some function $g(\cdot)$.

In general, $g(X)$ is not linear, i.e., not $a + bX$. It could be that $g(X) = a + bX + cX^2$.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite! Simple but most convenient.

Recall that $L[Y|X] = a + bX$ is a function of X .

This is similar: $E[Y|X] = g(X)$ for some function $g(\cdot)$.

In general, $g(X)$ is not linear, i.e., not $a + bX$. It could be that $g(X) = a + bX + cX^2$. Or that $g(X) = 2\sin(4X) + \exp\{-3X\}$.

Deja vu, all over again?

Have we seen this before? Yes.

Is anything new? Yes.

The idea of defining $g(x) = E[Y|X = x]$ and then $E[Y|X] = g(X)$.

Big deal? Quite! Simple but most convenient.

Recall that $L[Y|X] = a + bX$ is a function of X .

This is similar: $E[Y|X] = g(X)$ for some function $g(\cdot)$.

In general, $g(X)$ is not linear, i.e., not $a + bX$. It could be that $g(X) = a + bX + cX^2$. Or that $g(X) = 2\sin(4X) + \exp\{-3X\}$. Or something else.

Properties of CE

$$E[Y|X = x] = \sum_y yPr[Y = y|X = x]$$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

(a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

(a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;

(b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;

(c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;

(d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;

(e) $E[E[Y|X]] = E[Y]$.

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

(a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;

(b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;

(c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;

(d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;

(e) $E[E[Y|X]] = E[Y]$.

Proof:

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

(a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;

(b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;

(c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;

(d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;

(e) $E[E[Y|X]] = E[Y]$.

Proof:

(a),(b) Obvious

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

(a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;

(b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;

(c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;

(d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;

(e) $E[E[Y|X]] = E[Y]$.

Proof:

(a),(b) Obvious

(c) $E[Yh(X)|X = x] = \sum_{\omega} Y(\omega)h(X(\omega))\Pr[\omega|X = x]$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

(a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;

(b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;

(c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;

(d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;

(e) $E[E[Y|X]] = E[Y]$.

Proof:

(a),(b) Obvious

$$\begin{aligned} \text{(c)} \quad E[Yh(X)|X = x] &= \sum_{\omega} Y(\omega)h(X(\omega))\Pr[\omega|X = x] \\ &= \sum_{\omega} Y(\omega)h(x)\Pr[\omega|X = x] \end{aligned}$$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

(a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;

(b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;

(c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;

(d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;

(e) $E[E[Y|X]] = E[Y]$.

Proof:

(a),(b) Obvious

$$\begin{aligned} \text{(c)} \quad E[Yh(X)|X = x] &= \sum_{\omega} Y(\omega)h(X(\omega))\Pr[\omega|X = x] \\ &= \sum_{\omega} Y(\omega)h(x)\Pr[\omega|X = x] = h(x)E[Y|X = x] \end{aligned}$$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

$$(d) E[h(X)E[Y|X]] = \sum_x h(x)E[Y|X = x] \Pr[X = x]$$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

$$\begin{aligned} \text{(d) } E[h(X)E[Y|X]] &= \sum_x h(x) E[Y|X = x] \Pr[X = x] \\ &= \sum_x h(x) \sum_y y \Pr[Y = y|X = x] \Pr[X = x] \end{aligned}$$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

$$\begin{aligned} \text{(d) } E[h(X)E[Y|X]] &= \sum_x h(x) E[Y|X = x] \Pr[X = x] \\ &= \sum_x h(x) \sum_y y \Pr[Y = y|X = x] \Pr[X = x] \\ &= \sum_x h(x) \sum_y y \Pr[X = x, Y = y] \end{aligned}$$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

$$\begin{aligned} \text{(d) } E[h(X)E[Y|X]] &= \sum_x h(x) E[Y|X = x] \Pr[X = x] \\ &= \sum_x h(x) \sum_y y \Pr[Y = y|X = x] \Pr[X = x] \\ &= \sum_x h(x) \sum_y y \Pr[X = x, Y = y] \\ &= \sum_{x,y} h(x) y \Pr[X = x, Y = y] \end{aligned}$$

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

$$\begin{aligned} \text{(d) } E[h(X)E[Y|X]] &= \sum_x h(x) E[Y|X = x] \Pr[X = x] \\ &= \sum_x h(x) \sum_y y \Pr[Y = y|X = x] \Pr[X = x] \\ &= \sum_x h(x) \sum_y y \Pr[X = x, Y = y] \\ &= \sum_{x,y} h(x) y \Pr[X = x, Y = y] = E[h(X)Y]. \end{aligned}$$



Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

Properties of CE

$$E[Y|X = x] = \sum_y y \Pr[Y = y|X = x]$$

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Proof: (continued)

- (e) Let $h(X) = 1$ in (d).



Properties of CE

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Properties of CE

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Note that (d) says that

$$E[(Y - E[Y|X])h(X)] = 0.$$

Properties of CE

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Note that (d) says that

$$E[(Y - E[Y|X])h(X)] = 0.$$

We say that the estimation error $Y - E[Y|X]$ is orthogonal to every function $h(X)$ of X .

Properties of CE

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Note that (d) says that

$$E[(Y - E[Y|X])h(X)] = 0.$$

We say that the estimation error $Y - E[Y|X]$ is orthogonal to every function $h(X)$ of X .

We call this the **projection property**.

Properties of CE

Theorem

- (a) X, Y independent $\Rightarrow E[Y|X] = E[Y]$;
- (b) $E[aY + bZ|X] = aE[Y|X] + bE[Z|X]$;
- (c) $E[Yh(X)|X] = h(X)E[Y|X], \forall h(\cdot)$;
- (d) $E[h(X)E[Y|X]] = E[h(X)Y], \forall h(\cdot)$;
- (e) $E[E[Y|X]] = E[Y]$.

Note that (d) says that

$$E[(Y - E[Y|X])h(X)] = 0.$$

We say that the estimation error $Y - E[Y|X]$ is orthogonal to every function $h(X)$ of X .

We call this the **projection property**. More about this later.

Application: Calculating $E[Y|X]$

Let X, Y, Z be i.i.d. with mean 0 and variance 1.

Application: Calculating $E[Y|X]$

Let X, Y, Z be i.i.d. with mean 0 and variance 1. We want to calculate

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X].$$

Application: Calculating $E[Y|X]$

Let X, Y, Z be i.i.d. with mean 0 and variance 1. We want to calculate

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X].$$

We find

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X]$$

Application: Calculating $E[Y|X]$

Let X, Y, Z be i.i.d. with mean 0 and variance 1. We want to calculate

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X].$$

We find

$$\begin{aligned} E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X] \\ = 2 + 5X + 7XE[Y|X] + 11X^2 + 13X^3E[Z^2|X] \end{aligned}$$

Application: Calculating $E[Y|X]$

Let X, Y, Z be i.i.d. with mean 0 and variance 1. We want to calculate

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X].$$

We find

$$\begin{aligned} E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X] & \\ &= 2 + 5X + 7XE[Y|X] + 11X^2 + 13X^3E[Z^2|X] \\ &= 2 + 5X + 7XE[Y] + 11X^2 + 13X^3E[Z^2] \end{aligned}$$

Application: Calculating $E[Y|X]$

Let X, Y, Z be i.i.d. with mean 0 and variance 1. We want to calculate

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X].$$

We find

$$\begin{aligned} & E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X] \\ &= 2 + 5X + 7XE[Y|X] + 11X^2 + 13X^3E[Z^2|X] \\ &= 2 + 5X + 7XE[Y] + 11X^2 + 13X^3E[Z^2] \\ &= 2 + 5X + 11X^2 + 13X^3(\text{var}[Z] + E[Z]^2) \end{aligned}$$

Application: Calculating $E[Y|X]$

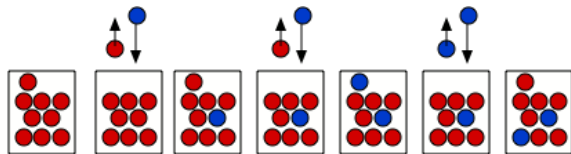
Let X, Y, Z be i.i.d. with mean 0 and variance 1. We want to calculate

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X].$$

We find

$$\begin{aligned} & E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X] \\ &= 2 + 5X + 7XE[Y|X] + 11X^2 + 13X^3E[Z^2|X] \\ &= 2 + 5X + 7XE[Y] + 11X^2 + 13X^3E[Z^2] \\ &= 2 + 5X + 11X^2 + 13X^3(\text{var}[Z] + E[Z]^2) \\ &= 2 + 5X + 11X^2 + 13X^3. \end{aligned}$$

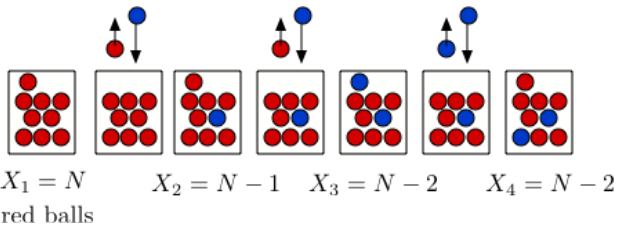
Application: Diluting



$X_1 = N$
red balls

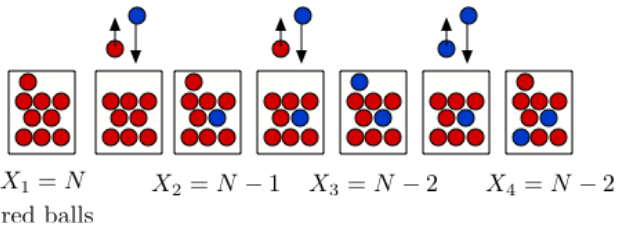
$X_2 = N - 1$ $X_3 = N - 2$ $X_4 = N - 2$

Application: Diluting



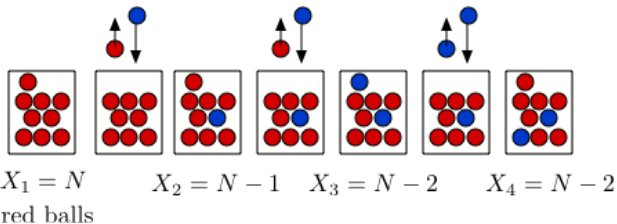
Each step, pick ball from well-mixed urn.

Application: Diluting



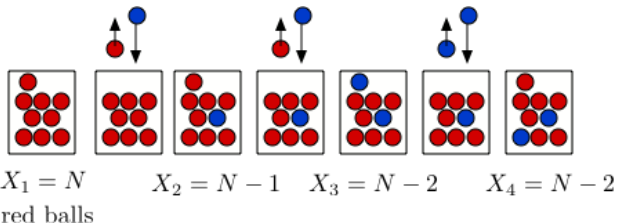
Each step, pick ball from well-mixed urn. Replace with blue ball.

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball.
Let X_n be the number of red balls in the urn at step n .

Application: Diluting

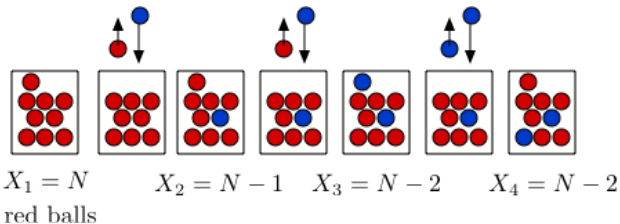


Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

Application: Diluting



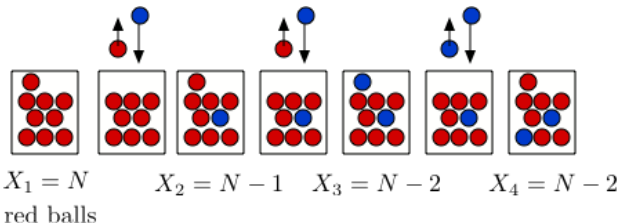
Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N

Application: Diluting



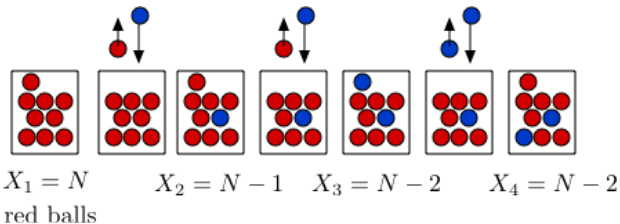
Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball)

Application: Diluting



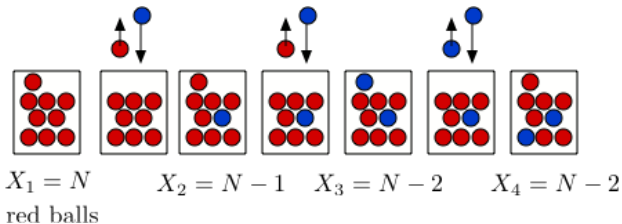
Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball)
and $X_{n+1} = m$ otherwise.

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball.

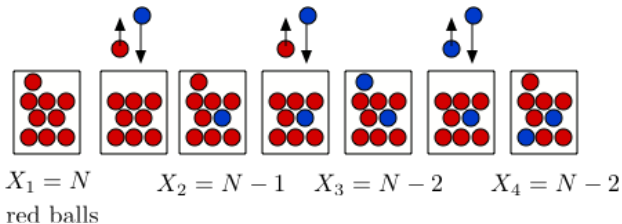
Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball)
and $X_{n+1} = m$ otherwise. Hence,

$$E[X_{n+1} | X_n = m] = m - (m/N)$$

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

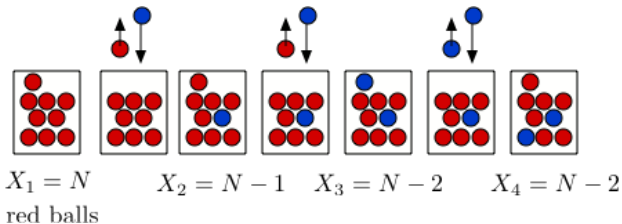
What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball) and $X_{n+1} = m$ otherwise. Hence,

$$E[X_{n+1} | X_n = m] = m - (m/N) = m(N - 1)/N = X_n \rho,$$

with $\rho := (N - 1)/N$.

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

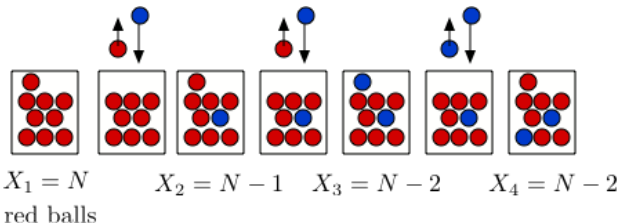
What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball) and $X_{n+1} = m$ otherwise. Hence,

$$E[X_{n+1} | X_n = m] = m - (m/N) = m(N - 1)/N = X_n \rho,$$

with $\rho := (N - 1)/N$. Consequently,

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

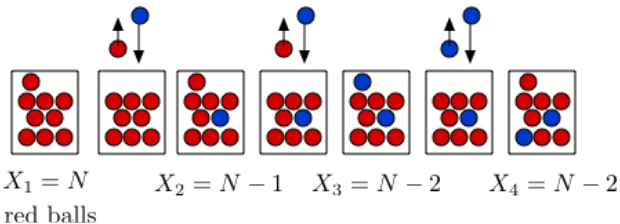
Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball) and $X_{n+1} = m$ otherwise. Hence,

$$E[X_{n+1} | X_n = m] = m - (m/N) = m(N - 1)/N = X_n \rho,$$

with $\rho := (N - 1)/N$. Consequently,

$$E[X_{n+1}] = E[E[X_{n+1} | X_n]]$$

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

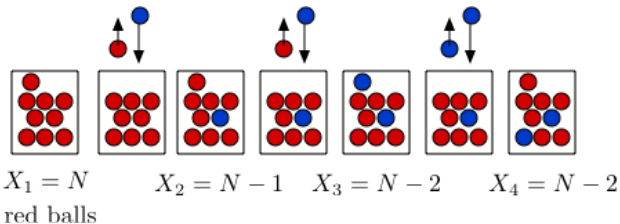
Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball) and $X_{n+1} = m$ otherwise. Hence,

$$E[X_{n+1} | X_n = m] = m - (m/N) = m(N - 1)/N = X_n \rho,$$

with $\rho := (N - 1)/N$. Consequently,

$$E[X_{n+1}] = E[E[X_{n+1} | X_n]] = \rho E[X_n], n \geq 1.$$

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball.

Let X_n be the number of red balls in the urn at step n .

What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball) and $X_{n+1} = m$ otherwise. Hence,

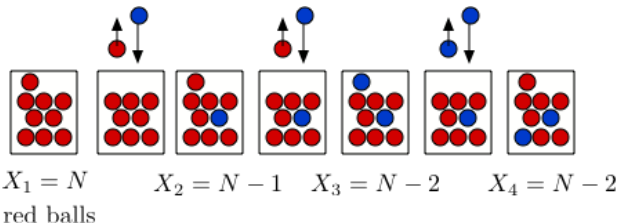
$$E[X_{n+1} | X_n = m] = m - (m/N) = m(N - 1)/N = X_n \rho,$$

with $\rho := (N - 1)/N$. Consequently,

$$E[X_{n+1}] = E[E[X_{n+1} | X_n]] = \rho E[X_n], n \geq 1.$$

$$\implies E[X_n] = \rho^{n-1} E[X_1]$$

Application: Diluting



Each step, pick ball from well-mixed urn. Replace with blue ball. Let X_n be the number of red balls in the urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball) and $X_{n+1} = m$ otherwise. Hence,

$$E[X_{n+1} | X_n = m] = m - (m/N) = m(N - 1)/N = X_n \rho,$$

with $\rho := (N - 1)/N$. Consequently,

$$E[X_{n+1}] = E[E[X_{n+1} | X_n]] = \rho E[X_n], n \geq 1.$$

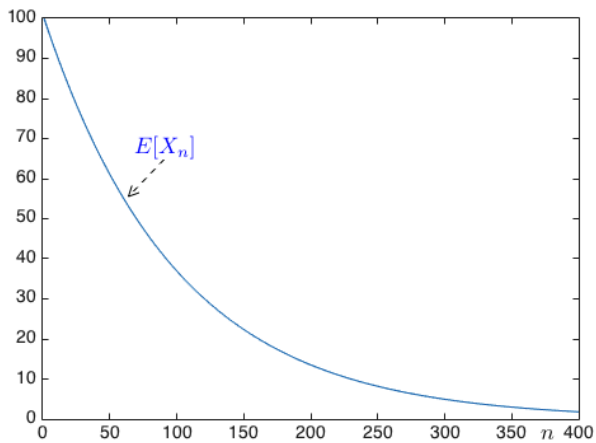
$$\implies E[X_n] = \rho^{n-1} E[X_1] = N \left(\frac{N-1}{N} \right)^{n-1}, n \geq 1.$$

Diluting

Here is a plot:

Diluting

Here is a plot:



Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N\left(\frac{N-1}{N}\right)^{n-1}, n \geq 1$.

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}, n \geq 1$.

Here is another argument for that result.

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N\left(\frac{N-1}{N}\right)^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked.

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}, n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$.

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}, n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$. Define:

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$. Define:

$$Y_n(k) = 1_{\{\text{ball } k \text{ is red at step } n\}}.$$

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$. Define:

$$Y_n(k) = 1\{\text{ball } k \text{ is red at step } n\}.$$

Then, $X_n = Y_n(1) + \dots + Y_n(N)$.

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$. Define:

$$Y_n(k) = 1\{\text{ball } k \text{ is red at step } n\}.$$

Then, $X_n = Y_n(1) + \dots + Y_n(N)$. Hence,

$$E[X_n] = E[Y_n(1) + \dots + Y_n(N)]$$

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$. Define:

$$Y_n(k) = 1\{\text{ball } k \text{ is red at step } n\}.$$

Then, $X_n = Y_n(1) + \dots + Y_n(N)$. Hence,

$$E[X_n] = E[Y_n(1) + \dots + Y_n(N)] = NE[Y_n(1)]$$

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$. Define:

$$Y_n(k) = 1 \{\text{ball } k \text{ is red at step } n\}.$$

Then, $X_n = Y_n(1) + \dots + Y_n(N)$. Hence,

$$\begin{aligned} E[X_n] &= E[Y_n(1) + \dots + Y_n(N)] = NE[Y_n(1)] \\ &= NPr[Y_n(1) = 1] \end{aligned}$$

Diluting

By analyzing $E[X_{n+1}|X_n]$, we found that $E[X_n] = N(\frac{N-1}{N})^{n-1}$, $n \geq 1$.

Here is another argument for that result.

Consider one particular red ball, say ball k .

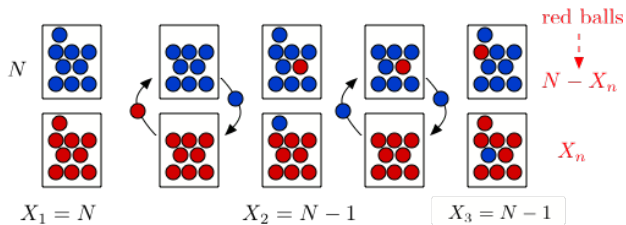
Each step, it remains red w.p. $(N-1)/N$, if different ball picked. \implies
the probability still red at step n is $[(N-1)/N]^{n-1}$. Define:

$$Y_n(k) = 1 \{\text{ball } k \text{ is red at step } n\}.$$

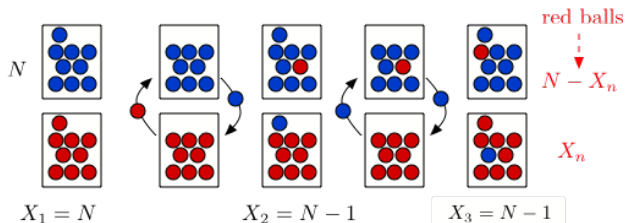
Then, $X_n = Y_n(1) + \dots + Y_n(N)$. Hence,

$$\begin{aligned} E[X_n] &= E[Y_n(1) + \dots + Y_n(N)] = NE[Y_n(1)] \\ &= NPr[Y_n(1) = 1] = N[(N-1)/N]^{n-1}. \end{aligned}$$

Application: Mixing

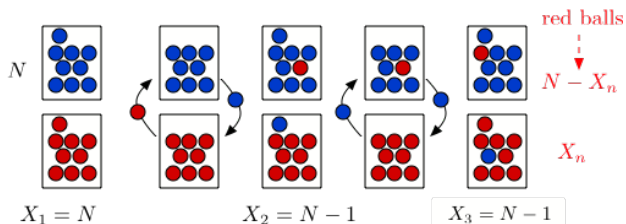


Application: Mixing



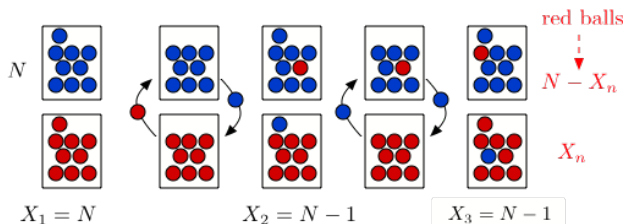
Each step, pick ball from each well-mixed urn.

Application: Mixing



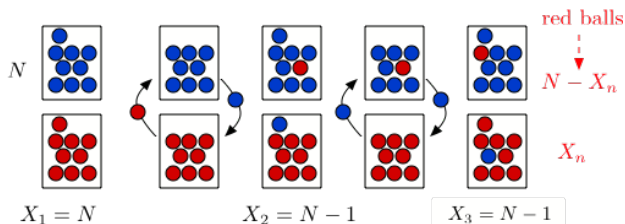
Each step, pick ball from each well-mixed urn. Transfer it to other urn.

Application: Mixing



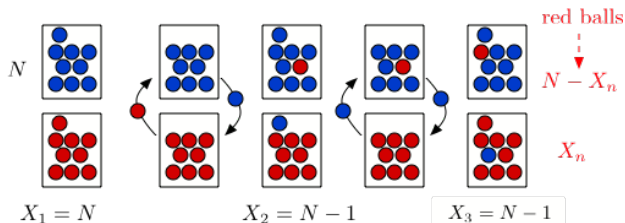
Each step, pick ball from each well-mixed urn. Transfer it to other urn. Let X_n be the number of red balls in the bottom urn at step n .

Application: Mixing



Each step, pick ball from each well-mixed urn. Transfer it to other urn.
Let X_n be the number of red balls in the bottom urn at step n .
What is $E[X_n]$?

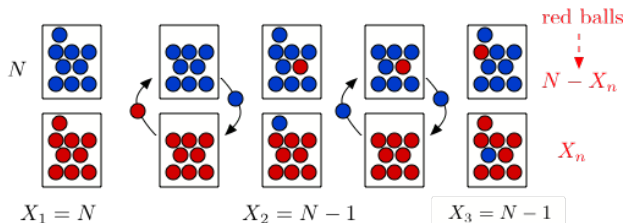
Application: Mixing



Each step, pick ball from each well-mixed urn. Transfer it to other urn. Let X_n be the number of red balls in the bottom urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m + 1$ w.p. p and $X_{n+1} = m - 1$ w.p. q

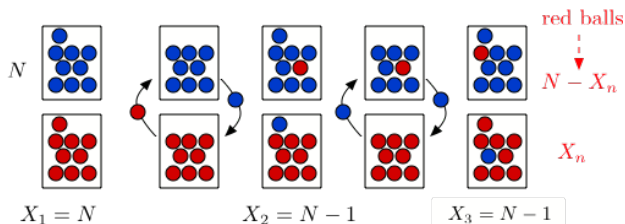
Application: Mixing



Each step, pick ball from each well-mixed urn. Transfer it to other urn. Let X_n be the number of red balls in the bottom urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m + 1$ w.p. p and $X_{n+1} = m - 1$ w.p. q where $p = (1 - m/N)^2$ (B goes up, R down)

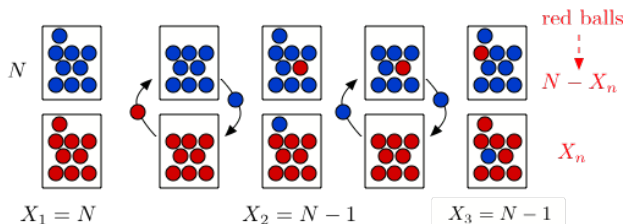
Application: Mixing



Each step, pick ball from each well-mixed urn. Transfer it to other urn. Let X_n be the number of red balls in the bottom urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m + 1$ w.p. p and $X_{n+1} = m - 1$ w.p. q where $p = (1 - m/N)^2$ (B goes up, R down) and $q = (m/N)^2$ (R goes up, B down).

Application: Mixing



Each step, pick ball from each well-mixed urn. Transfer it to other urn.

Let X_n be the number of red balls in the bottom urn at step n .

What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m + 1$ w.p. p and $X_{n+1} = m - 1$ w.p. q

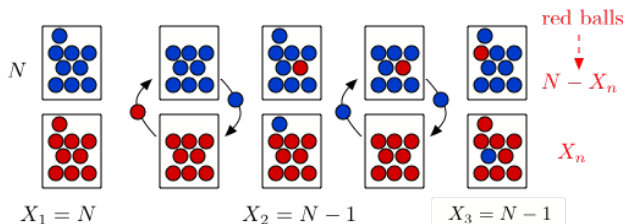
where $p = (1 - m/N)^2$ (B goes up, R down)

and $q = (m/N)^2$ (R goes up, B down).

Thus,

$$E[X_{n+1}|X_n] = X_n + p - q$$

Application: Mixing



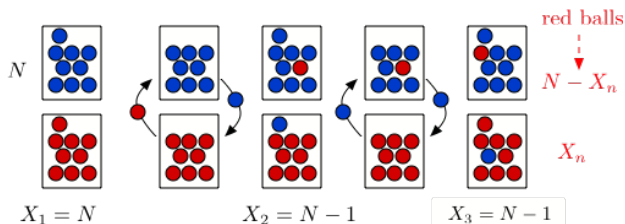
Each step, pick ball from each well-mixed urn. Transfer it to other urn. Let X_n be the number of red balls in the bottom urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m + 1$ w.p. p and $X_{n+1} = m - 1$ w.p. q where $p = (1 - m/N)^2$ (B goes up, R down) and $q = (m/N)^2$ (R goes up, B down).

Thus,

$$E[X_{n+1}|X_n] = X_n + p - q = X_n + 1 - 2X_n/N$$

Application: Mixing



Each step, pick ball from each well-mixed urn. Transfer it to other urn. Let X_n be the number of red balls in the bottom urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m + 1$ w.p. p and $X_{n+1} = m - 1$ w.p. q
where $p = (1 - m/N)^2$ (B goes up, R down)
and $q = (m/N)^2$ (R goes up, B down).

Thus,

$$E[X_{n+1}|X_n] = X_n + p - q = X_n + 1 - 2X_n/N = 1 + \rho X_n, \quad \rho := (1 - 2/N).$$

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$.

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$.

Does that make sense?

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$.

Does that make sense?

Hence,

$$E[X_{n+1}] = 1 + \rho E[X_n]$$

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$.

Does that make sense?

Hence,

$$E[X_{n+1}] = 1 + \rho E[X_n]$$

$$E[X_2] = 1 + \rho N; E[X_3] = 1 + \rho(1 + \rho N) = 1 + \rho + \rho^2 N$$

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$.

Does that make sense?

Hence,

$$E[X_{n+1}] = 1 + \rho E[X_n]$$

$$E[X_2] = 1 + \rho N; E[X_3] = 1 + \rho(1 + \rho N) = 1 + \rho + \rho^2 N$$

$$E[X_4] = 1 + \rho(1 + \rho + \rho^2 N) = 1 + \rho + \rho^2 + \rho^3 N$$

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$.

Does that make sense?

Hence,

$$E[X_{n+1}] = 1 + \rho E[X_n]$$

$$E[X_2] = 1 + \rho N; E[X_3] = 1 + \rho(1 + \rho N) = 1 + \rho + \rho^2 N$$

$$E[X_4] = 1 + \rho(1 + \rho + \rho^2 N) = 1 + \rho + \rho^2 + \rho^3 N$$

$$E[X_n] = 1 + \rho + \dots + \rho^{n-2} + \rho^{n-1} N.$$

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$.

Does that make sense?

Hence,

$$E[X_{n+1}] = 1 + \rho E[X_n]$$

$$E[X_2] = 1 + \rho N; E[X_3] = 1 + \rho(1 + \rho N) = 1 + \rho + \rho^2 N$$

$$E[X_4] = 1 + \rho(1 + \rho + \rho^2 N) = 1 + \rho + \rho^2 + \rho^3 N$$

$$E[X_n] = 1 + \rho + \dots + \rho^{n-2} + \rho^{n-1} N.$$

Hence,

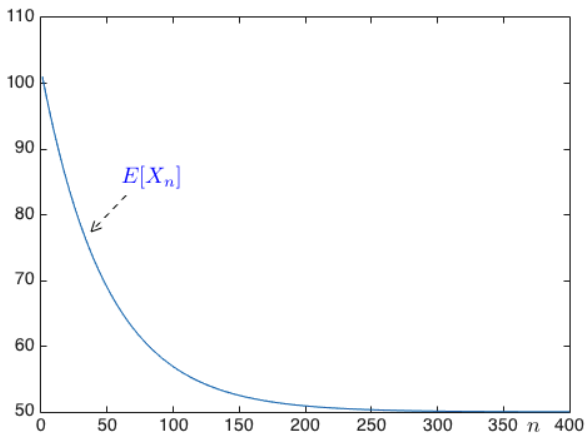
$$E[X_n] = \frac{1 - \rho^{n-1}}{1 - \rho} + \rho^{n-1} N, n \geq 1.$$

Application: Mixing

Here is the plot.

Application: Mixing

Here is the plot.



Application: Going Viral

Consider a social network (e.g., Twitter).

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends.

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends. Each of your friend retweets w.p. p .

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends. Each of your friend retweets w.p. p .

Each of your friends has d friends, etc.

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends. Each of your friend retweets w.p. p .

Each of your friends has d friends, etc.

Does the rumor spread?

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends. Each of your friend retweets w.p. p .

Each of your friends has d friends, etc.

Does the rumor spread? Does it die out

Application: Going Viral

Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends. Each of your friend retweets w.p. p .

Each of your friends has d friends, etc.

Does the rumor spread? Does it die out (mercifully)?

Application: Going Viral

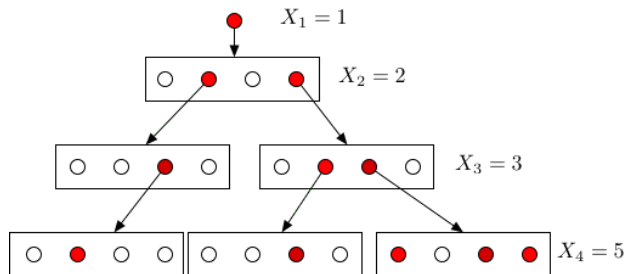
Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends. Each of your friend retweets w.p. p .

Each of your friends has d friends, etc.

Does the rumor spread? Does it die out (mercifully)?



Application: Going Viral

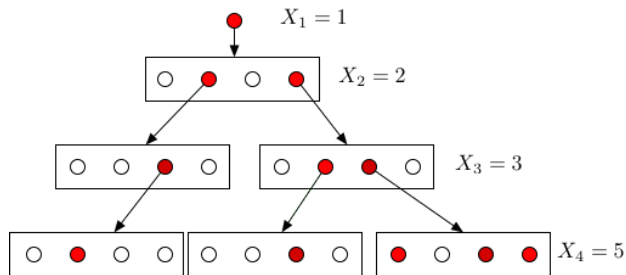
Consider a social network (e.g., Twitter).

You start a rumor (e.g., Rao is bad at making copies).

You have d friends. Each of your friend retweets w.p. p .

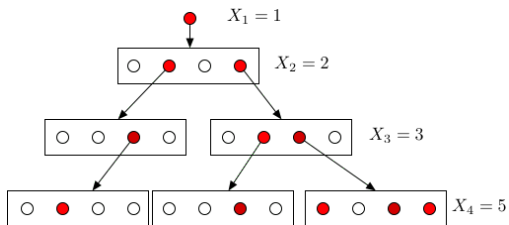
Each of your friends has d friends, etc.

Does the rumor spread? Does it die out (mercifully)?

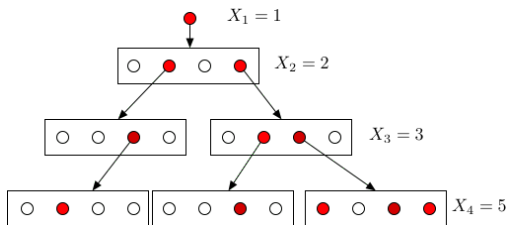


In this example, $d = 4$.

Application: Going Viral

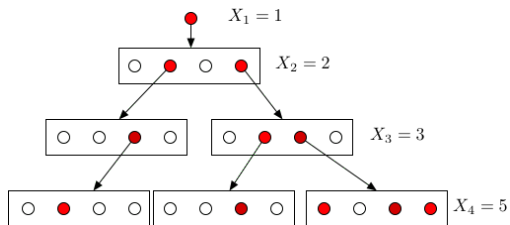


Application: Going Viral



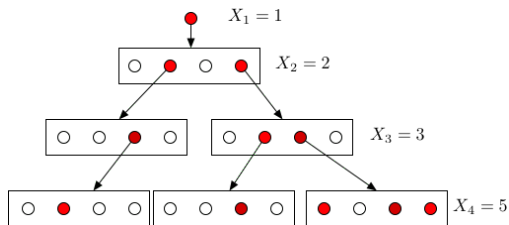
Fact:

Application: Going Viral



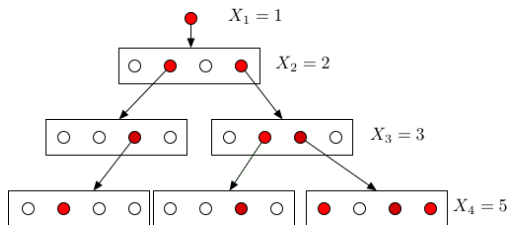
Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n .

Application: Going Viral



Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n .
Then, $E[X] < \infty$ iff $pd < 1$.

Application: Going Viral

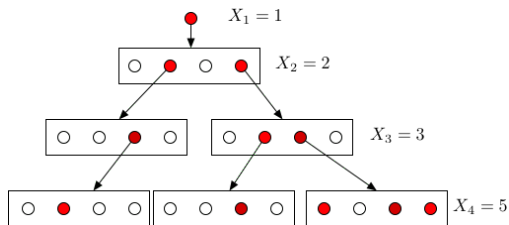


Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n .
Then, $E[X] < \infty$ iff $pd < 1$.

Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$.

Application: Going Viral

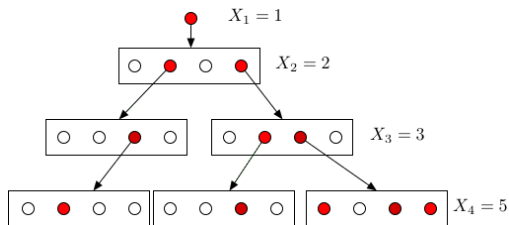


Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n . Then, $E[X] < \infty$ iff $pd < 1$.

Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1} | X_n = k] = kpd$.

Application: Going Viral



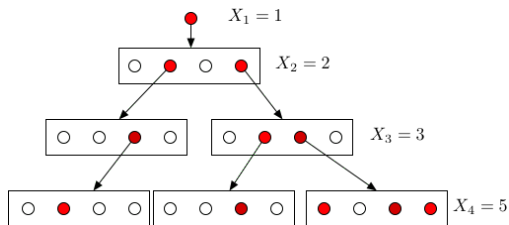
Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n . Then, $E[X] < \infty$ iff $pd < 1$.

Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1} | X_n = k] = kpd$.

Thus, $E[X_{n+1} | X_n] = pdX_n$.

Application: Going Viral



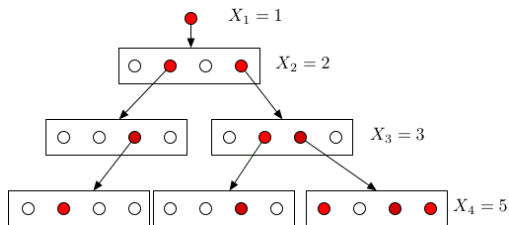
Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n . Then, $E[X] < \infty$ iff $pd < 1$.

Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1} | X_n = k] = kpd$.

Thus, $E[X_{n+1} | X_n] = pdX_n$. Consequently, $E[X_n] = (pd)^{n-1}$, $n \geq 1$.

Application: Going Viral



Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n . Then, $E[X] < \infty$ iff $pd < 1$.

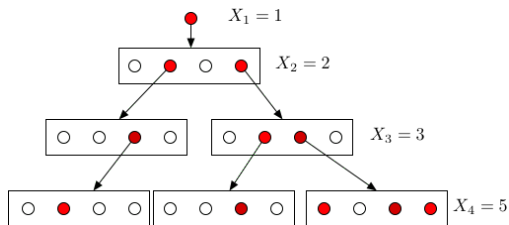
Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1} | X_n = k] = kpd$.

Thus, $E[X_{n+1} | X_n] = pdX_n$. Consequently, $E[X_n] = (pd)^{n-1}$, $n \geq 1$.

If $pd < 1$, then $E[X_1 + \dots + X_n] \leq (1 - pd)^{-1}$

Application: Going Viral



Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n . Then, $E[X] < \infty$ iff $pd < 1$.

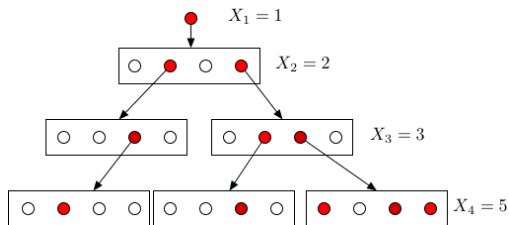
Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1} | X_n = k] = kpd$.

Thus, $E[X_{n+1} | X_n] = pdX_n$. Consequently, $E[X_n] = (pd)^{n-1}$, $n \geq 1$.

If $pd < 1$, then $E[X_1 + \dots + X_n] \leq (1 - pd)^{-1} \implies E[X] \leq (1 - pd)^{-1}$.

Application: Going Viral



Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n . Then, $E[X] < \infty$ iff $pd < 1$.

Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1} | X_n = k] = kpd$.

Thus, $E[X_{n+1} | X_n] = pdX_n$. Consequently, $E[X_n] = (pd)^{n-1}$, $n \geq 1$.

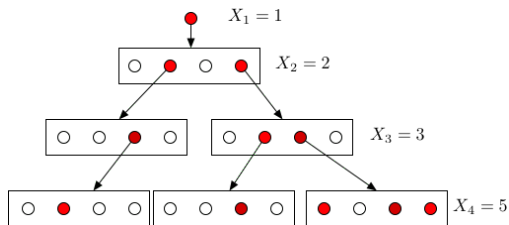
If $pd < 1$, then $E[X_1 + \dots + X_n] \leq (1 - pd)^{-1} \implies E[X] \leq (1 - pd)^{-1}$.

If $pd \geq 1$, then for all C one can find n s.t.

$$E[X] \geq E[X_1 + \dots + X_n] \geq C.$$



Application: Going Viral



Fact: Number of tweets $X = \sum_{n=1}^{\infty} X_n$ where X_n is tweets in level n . Then, $E[X] < \infty$ iff $pd < 1$.

Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1}|X_n = k] = kpd$.

Thus, $E[X_{n+1}|X_n] = pdX_n$. Consequently, $E[X_n] = (pd)^{n-1}$, $n \geq 1$.

If $pd < 1$, then $E[X_1 + \dots + X_n] \leq (1 - pd)^{-1} \implies E[X] \leq (1 - pd)^{-1}$.

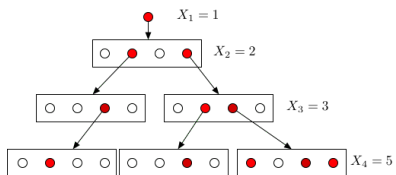
If $pd \geq 1$, then for all C one can find n s.t.

$$E[X] \geq E[X_1 + \dots + X_n] \geq C.$$

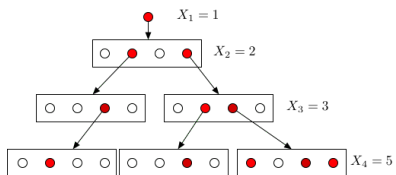
□

In fact, one can show that $pd \geq 1 \implies Pr[X = \infty] > 0$.

Application: Going Viral

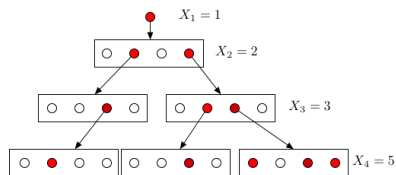


Application: Going Viral



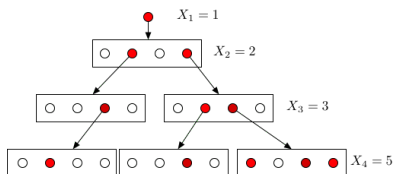
An easy extension:

Application: Going Viral



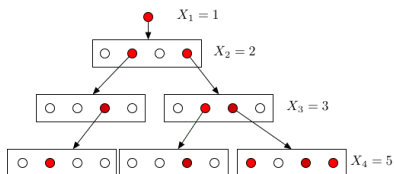
An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$.

Application: Going Viral



An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

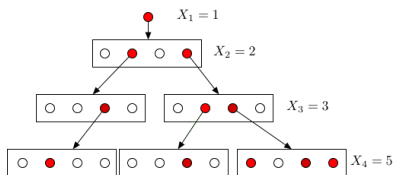
Application: Going Viral



An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people,

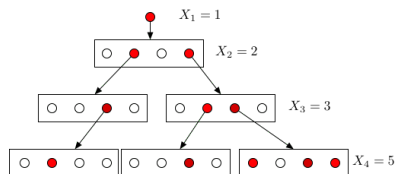
Application: Going Viral



An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people, one has $X_{n+1} = B(d_1 + \dots + d_k, p)$.

Application: Going Viral

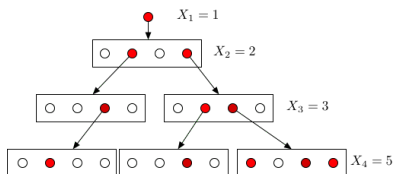


An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people, one has $X_{n+1} = B(d_1 + \dots + d_k, p)$. Hence,

$$E[X_{n+1} | X_n = k, D_1 = d_1, \dots, D_k = d_k] = p(d_1 + \dots + d_k).$$

Application: Going Viral



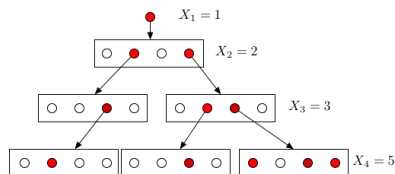
An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people, one has $X_{n+1} = B(d_1 + \dots + d_k, p)$. Hence,

$$E[X_{n+1} | X_n = k, D_1 = d_1, \dots, D_k = d_k] = p(d_1 + \dots + d_k).$$

Thus, $E[X_{n+1} | X_n = k, D_1, \dots, D_k] = p(D_1 + \dots + D_k)$.

Application: Going Viral



An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

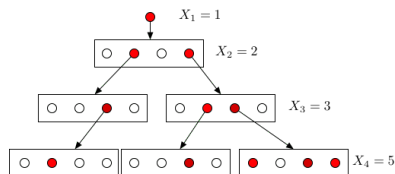
To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people, one has $X_{n+1} = B(d_1 + \dots + d_k, p)$. Hence,

$$E[X_{n+1} | X_n = k, D_1 = d_1, \dots, D_k = d_k] = p(d_1 + \dots + d_k).$$

Thus, $E[X_{n+1} | X_n = k, D_1, \dots, D_k] = p(D_1 + \dots + D_k)$.

Consequently, $E[X_{n+1} | X_n = k] = E[p(D_1 + \dots + D_k)] = pdk$.

Application: Going Viral



An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people, one has $X_{n+1} = B(d_1 + \dots + d_k, p)$. Hence,

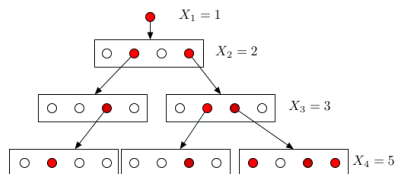
$$E[X_{n+1} | X_n = k, D_1 = d_1, \dots, D_k = d_k] = p(d_1 + \dots + d_k).$$

Thus, $E[X_{n+1} | X_n = k, D_1, \dots, D_k] = p(D_1 + \dots + D_k)$.

Consequently, $E[X_{n+1} | X_n = k] = E[p(D_1 + \dots + D_k)] = pdk$.

Finally, $E[X_{n+1} | X_n] = pdX_n$,

Application: Going Viral



An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people, one has $X_{n+1} = B(d_1 + \dots + d_k, p)$. Hence,

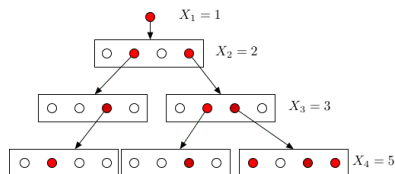
$$E[X_{n+1} | X_n = k, D_1 = d_1, \dots, D_k = d_k] = p(d_1 + \dots + d_k).$$

Thus, $E[X_{n+1} | X_n = k, D_1, \dots, D_k] = p(D_1 + \dots + D_k)$.

Consequently, $E[X_{n+1} | X_n = k] = E[p(D_1 + \dots + D_k)] = pdk$.

Finally, $E[X_{n+1} | X_n] = pdX_n$, and $E[X_{n+1}] = pdE[X_n]$.

Application: Going Viral



An easy extension: Assume that everyone has an independent number D_i of friends with $E[D_i] = d$. Then, the same fact holds.

To see this, note that given $X_n = k$, and given the numbers of friends $D_1 = d_1, \dots, D_k = d_k$ of these X_n people, one has $X_{n+1} = B(d_1 + \dots + d_k, p)$. Hence,

$$E[X_{n+1} | X_n = k, D_1 = d_1, \dots, D_k = d_k] = p(d_1 + \dots + d_k).$$

Thus, $E[X_{n+1} | X_n = k, D_1, \dots, D_k] = p(D_1 + \dots + D_k)$.

Consequently, $E[X_{n+1} | X_n = k] = E[p(D_1 + \dots + D_k)] = pdk$.

Finally, $E[X_{n+1} | X_n] = pdX_n$, and $E[X_{n+1}] = pdE[X_n]$.

We conclude as before.

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

and $E[X_n] = \mu$ for all $n \geq 1$.

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

and $E[X_n] = \mu$ for all $n \geq 1$.

Then,

$$E[X_1 + \dots + X_Z] = \mu E[Z].$$

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

and $E[X_n] = \mu$ for all $n \geq 1$.

Then,

$$E[X_1 + \dots + X_Z] = \mu E[Z].$$

Proof:

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

and $E[X_n] = \mu$ for all $n \geq 1$.

Then,

$$E[X_1 + \dots + X_Z] = \mu E[Z].$$

Proof:

$$E[X_1 + \dots + X_Z | Z = k] = \mu k.$$

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

and $E[X_n] = \mu$ for all $n \geq 1$.

Then,

$$E[X_1 + \dots + X_Z] = \mu E[Z].$$

Proof:

$$E[X_1 + \dots + X_Z | Z = k] = \mu k.$$

$$\text{Thus, } E[X_1 + \dots + X_Z | Z] = \mu Z.$$

Application: Wald's Identity

Here is an extension of an identity we used in the last slide.

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

and $E[X_n] = \mu$ for all $n \geq 1$.

Then,

$$E[X_1 + \dots + X_Z] = \mu E[Z].$$

Proof:

$$E[X_1 + \dots + X_Z | Z = k] = \mu k.$$

$$\text{Thus, } E[X_1 + \dots + X_Z | Z] = \mu Z.$$

$$\text{Hence, } E[X_1 + \dots + X_Z] = E[\mu Z] = \mu E[Z].$$



CE = MMSE

Theorem

$E[Y|X]$ is the 'best' guess about Y based on X .

CE = MMSE

Theorem

$E[Y|X]$ is the 'best' guess about Y based on X .

Specifically, it is the function $g(X)$ of X that

minimizes $E[(Y - g(X))^2]$.

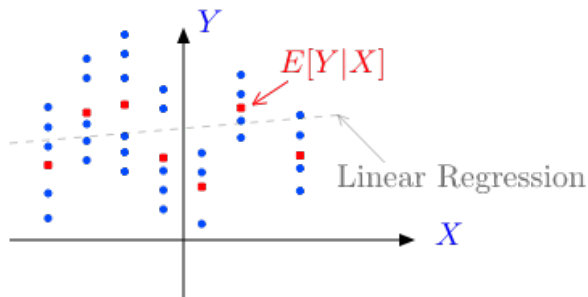
CE = MMSE

Theorem

$E[Y|X]$ is the 'best' guess about Y based on X .

Specifically, it is the function $g(X)$ of X that

minimizes $E[(Y - g(X))^2]$.



CE = MMSE

Theorem CE = MMSE

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

Let $h(X)$ be any function of X .

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

Let $h(X)$ be any function of X . Then

$$E[(Y - h(X))^2] =$$

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

Let $h(X)$ be any function of X . Then

$$E[(Y - h(X))^2] = E[(Y - g(X) + g(X) - h(X))^2]$$

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

Let $h(X)$ be any function of X . Then

$$\begin{aligned} E[(Y - h(X))^2] &= E[(Y - g(X) + g(X) - h(X))^2] \\ &= E[(Y - g(X))^2] + E[(g(X) - h(X))^2] \\ &\quad + 2E[(Y - g(X))(g(X) - h(X))]. \end{aligned}$$

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

Let $h(X)$ be any function of X . Then

$$\begin{aligned} E[(Y - h(X))^2] &= E[(Y - g(X) + g(X) - h(X))^2] \\ &= E[(Y - g(X))^2] + E[(g(X) - h(X))^2] \\ &\quad + 2E[(Y - g(X))(g(X) - h(X))]. \end{aligned}$$

But,

$$E[(Y - g(X))(g(X) - h(X))] = 0 \text{ by the projection property.}$$

CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

Let $h(X)$ be any function of X . Then

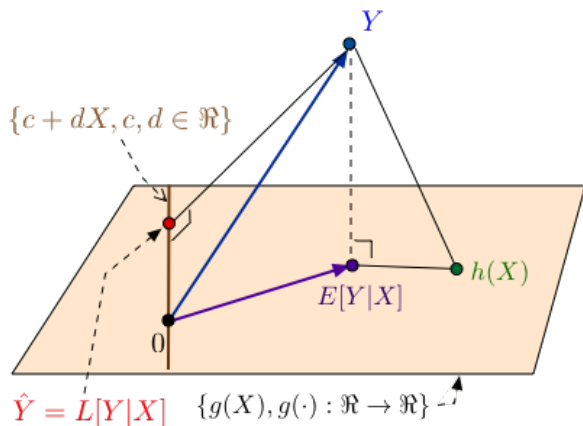
$$\begin{aligned} E[(Y - h(X))^2] &= E[(Y - g(X) + g(X) - h(X))^2] \\ &= E[(Y - g(X))^2] + E[(g(X) - h(X))^2] \\ &\quad + 2E[(Y - g(X))(g(X) - h(X))]. \end{aligned}$$

But,

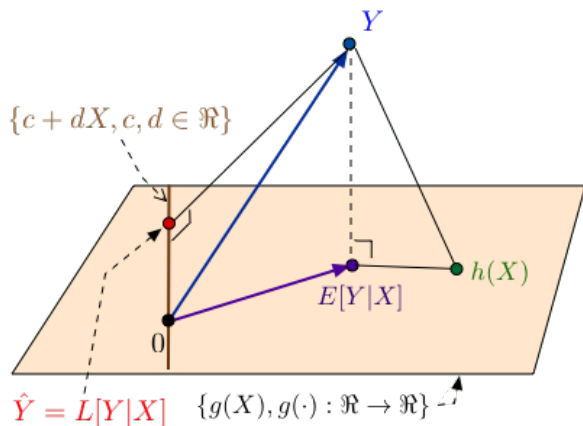
$$E[(Y - g(X))(g(X) - h(X))] = 0 \text{ by the projection property.}$$

Thus, $E[(Y - h(X))^2] \geq E[(Y - g(X))^2]$. □

$E[Y|X]$ and $L[Y|X]$ as projections

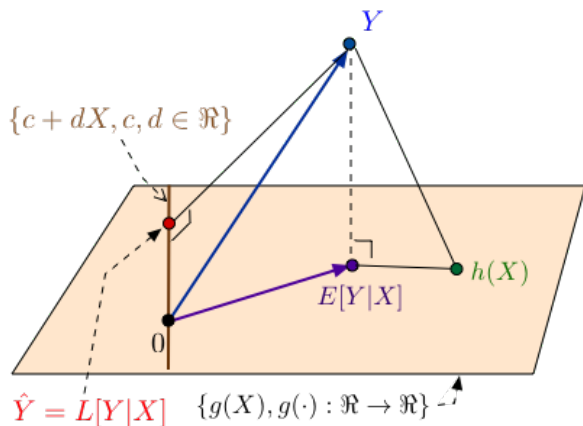


$E[Y|X]$ and $L[Y|X]$ as projections



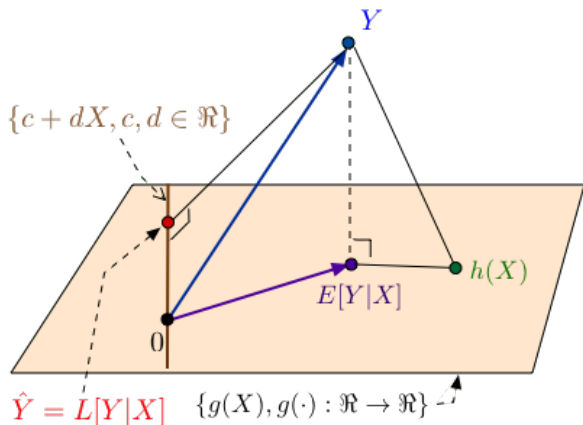
$L[Y|X]$ is the projection of Y on $\{a + bX, a, b \in \mathbb{R}\}$:

$E[Y|X]$ and $L[Y|X]$ as projections



$L[Y|X]$ is the projection of Y on $\{a + bX, a, b \in \mathbb{R}\}$: LLSE

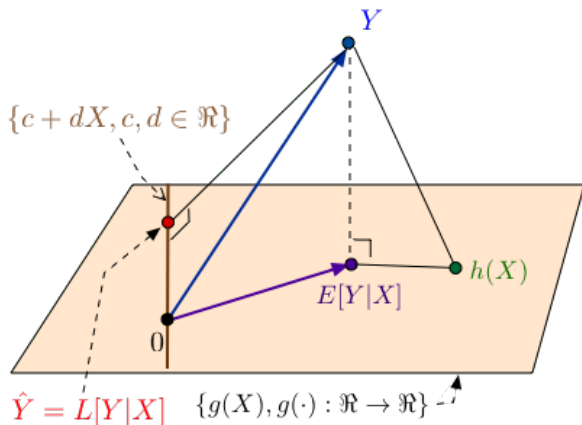
$E[Y|X]$ and $L[Y|X]$ as projections



$L[Y|X]$ is the projection of Y on $\{a + bX, a, b \in \mathbb{R}\}$: LLSE

$E[Y|X]$ is the projection of Y on $\{g(X), g(\cdot) : \mathbb{R} \rightarrow \mathbb{R}\}$:

$E[Y|X]$ and $L[Y|X]$ as projections



$L[Y|X]$ is the projection of Y on $\{a + bX, a, b \in \mathbb{R}\}$: LLSE

$E[Y|X]$ is the projection of Y on $\{g(X), g(\cdot) : \mathbb{R} \rightarrow \mathbb{R}\}$: MMSE.

Summary

Conditional Expectation

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$;

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:
 - ▶ Calculating $E[Y|X]$

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:
 - ▶ Calculating $E[Y|X]$
 - ▶ Diluting

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:
 - ▶ Calculating $E[Y|X]$
 - ▶ Diluting
 - ▶ Mixing

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:
 - ▶ Calculating $E[Y|X]$
 - ▶ Diluting
 - ▶ Mixing
 - ▶ Rumors

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:
 - ▶ Calculating $E[Y|X]$
 - ▶ Diluting
 - ▶ Mixing
 - ▶ Rumors
 - ▶ Wald

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity, $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:
 - ▶ Calculating $E[Y|X]$
 - ▶ Diluting
 - ▶ Mixing
 - ▶ Rumors
 - ▶ Wald
- ▶ MMSE: $E[Y|X]$ minimizes $E[(Y - g(X))^2]$ over all $g(\cdot)$